

## Solution for exercise 5.4.19 in Pitman

We apply the technique of the proof for the distribution of ratios formula page 382-383. Define  $Z = \frac{X}{X+Y}$ . The event  $z < Z < z + dz$  occurs whenever  $Y$  is between the two lines  $\frac{x}{x+y} = z + dz$  and  $\frac{x}{x+y} = z$ . We get the length of the vertical side of the rectangle by solving for  $y$  in the two equations above. Thus

$$y_2 = \left(\frac{1}{z} - 1\right)x, y_1 = \left(\frac{1}{z+dz} - 1\right)x, \quad y_2 - y_1 = \frac{xdz}{z(z+dz)} \approx \frac{xdz}{z^2}$$

We have derived a general formula for the density of  $Z = \frac{X}{X+Y}$  for non negative  $X$  and  $Y$

$$\int_0^\infty \frac{x}{z^2} f_X(x) f_Y\left(\frac{(1-z)x}{z}\right) dx$$

We now insert the gamma densities of  $X$  and  $Y$  (see page 481) to get

$$\int_0^\infty \frac{x}{z^2} \lambda \frac{(\lambda x)^{r-1}}{\Gamma(r)} e^{-\lambda x} \lambda \frac{\left(\lambda \frac{(1-z)x}{z}\right)^{s-1}}{\Gamma(s)} e^{-\lambda \frac{(1-z)x}{z}} dx$$

We simplify to get

$$\frac{1}{z^2 \Gamma(r) \Gamma(s)} \left(\frac{1-z}{z}\right)^{s-1} \int_0^\infty \lambda (\lambda x)^{r+s-1} e^{-\lambda \frac{x}{z}} dx$$

the function under the integral is very close to a gamma density such that with

$$\frac{1}{z^2 \Gamma(r) \Gamma(s)} \left(\frac{1-z}{z}\right)^{s-1} \Gamma(r+s) z^{r+s-1} \int_0^\infty \lambda \frac{\left(\lambda \frac{x}{z}\right)^{r+s-1}}{\Gamma(r+s)} e^{-\lambda \frac{x}{z}} dx$$

we get the density of a *gamma*  $(r+s, \frac{\lambda}{z})$  variable. Thus

$$f_Z(z) = \frac{\Gamma(r+s)}{\Gamma(r)\Gamma(s)} (1-z)^{s-1} z^{r+s-(s-1)-2} = \frac{1}{B(r,s)} z^{r-1} (1-z)^{s-1}$$

the density of a *beta*( $r, s$ ) random variable.

## Independence

Three lines to follow

1. We see directly from the calculations

2. Considering

$$P\left(z < \frac{X}{X+Y} < z + dz \mid w < X+Y < w + dw\right) = \frac{P\left(z < \frac{X}{X+Y} < z + dz, w < X+Y < w + dw\right)}{P(w < X+Y < w + dw)}$$

3. Using the division rule page 425